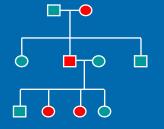
Family Based Association Studies

Gonçalo Abecasis University of Michigan

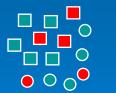
Design of Gene Mapping Studies How to find disease susceptibility alleles?

Magnitude of effect



Rare, high penetrance mutations – use linkage

Common, low penetrance variants – use association



Frequency in population



Genetic Linkage Studies

- Search for long stretches of chromosome shared between close relatives with similar phenotype
- Identify variants with relatively large contributions to disease risk
- Require only a coarse measurement of genetic variation
 - 400 800 microsatellites or a few thousand SNPs extract most of the linkage information in typical pedigrees
- High-throughput SNP genotyping has already sped up and facilitated these studies
 - Data analysis methods must select subset of independent SNPs or model disequilibrium between markers

Genetic Association Studies

Search for short stretches of chromosome shared between distantly related individuals with similar phenotype

Identify genetic variants with relatively small individual contributions to disease risk

Require detailed measurement of genetic variation

- >8,000,000 catalogued genetic variants, so ...
- Until recently, limited to candidate genes or regions
 - A hit-and-miss approach...

SNP resources and decreasing assay costs now make it possible to conduct comprehensive genome-wide scans

Association Studies in Families

Majority of association studies genotype unrelated individuals and nearly all linkage studies genotype related individuals...

However, tests for genetic association can use family data when relatedness between individuals is modeled appropriately (e.g. George and Elston, 1987)

Linear Model for Association

Model association using a model such as:

$$E(y_i) = \mu + \beta_g g + \beta_c c + \dots$$

- > y_i is the phenotype for individual *i*
- > g_i is the phenotype for individual *i*
 - Simplest coding is to set g_i = number of copies of allele '1'
- \succ c_i is a covariate for individual *i*
 - Covariates could be estimated ancestry, environmental factors...

 β coefficients are estimated effects of genotype, covariates

Allowing for Related Data

Similarities between individuals
 Variance–covariance matrix
 Major gene, polygenes, environment

 $\Omega_{ij} = \begin{cases} \sigma_a^2 + \sigma_g^2 + \sigma_e^2 & i = j \\ 2\varphi_{marker(ij)}\sigma_a^2 + 2\varphi_{ij}\sigma_g^2 & i \neq j \end{cases}$

Where,

 φ_{ij} is the kinship coefficient for the individuals i and j $\varphi_{marker(ij)}$ depends on the number of alleles shared IBD

Disequilibrium Mapping

Control for possible population structure

- Distinguish linkage disequilibrium from other types of association
- Family-based association analysis
 - Using families collected for linkage mapping
- Powerful if assumptions are met
 - Same disease haplotype shared by many patients
- > High-resolution

Controlling for Stratification

> If stratum were known...

- For each individual genotype (g_{ii})
- Average number of alleles in a strata (b_{ii})
- Adjust for stratum differences $(w_{ij} = g_{ij} b_{ij})$

$$\hat{y}_{ij} = \mu + \hat{\beta}_b b_{ij} + \hat{\beta}_w w_{ij}$$

How to define stratum then?
 Use family data to estimate b_{ii}

b_{ii} as Family Control

Expected genotype for each individual

- Ancestors
- Siblings

Informative individuals

- Genotype may differ from expected
- Have heterozygous ancestor in pedigree

Nuclear Families

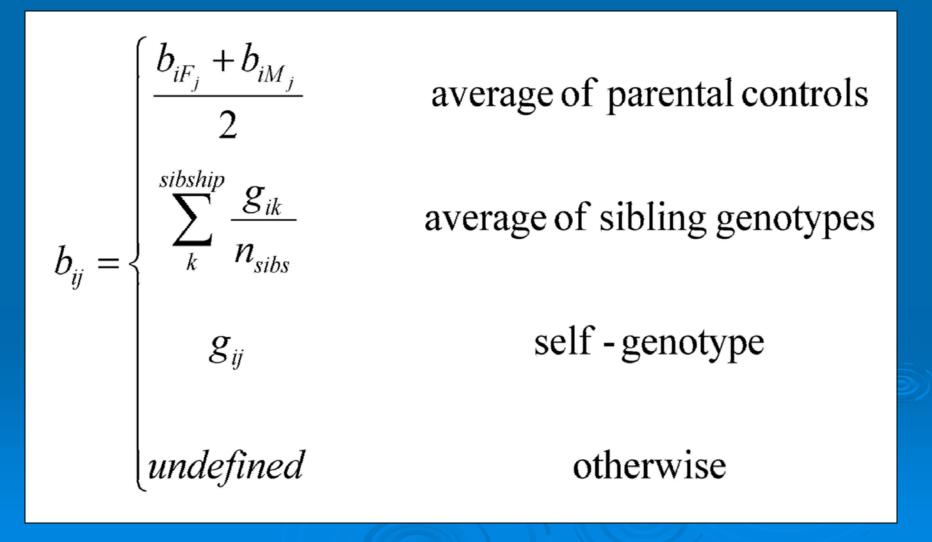
$$b_{ij} = b_i = \begin{cases} \frac{g_{iF} + g_{iM}}{2} & \text{average} \\ \frac{sibship}{\sum_{k} \frac{g_{ik}}{n_{sibs}}} & \text{average} \end{cases}$$

average of parental genotypes

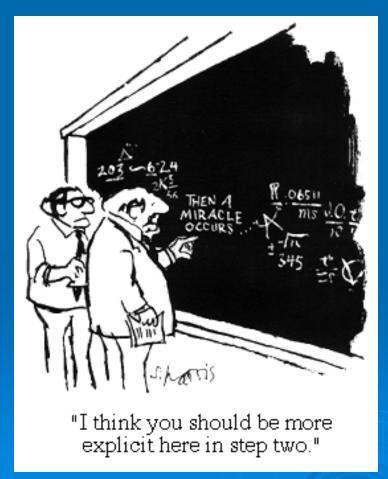
average of sibling genotypes.

$$w_{ij} = g_{ij} - b_{ij}$$

Extended Families



Parameter Derivations



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 $Model = (\mu, \beta_b, \beta_w, \sigma_e^2, \sigma_g^2, \sigma_a^2)$ $\begin{bmatrix} \beta_b \\ \vdots \\ \beta_w \end{bmatrix} = \begin{bmatrix} \frac{\sum_i n_i (p_i - q_i) \mu_i}{NV_b} + a \end{bmatrix}$ a $a = \frac{D}{pq} a_{QTL}$ $\sigma_a^2 = V_{OTL} - 2pqa$

QTDT Practical

> We'll use QTDT to:

- Evaluate whether a trait is heritable
- Evaluate evidence for linkage
- Evaluate evidence for association

We'll use the Bernard Keavney's classic ACE data as an example...

Application: Angiotensin-1

British population Circulating ACE levels Normalized separately for males / females > 10 di-allelic polymorphisms • 26 kb Common In strong linkage disequilibrium Keavney et al, HMG, 1998 goncaloa\keavney

Variance Components (QTDT)

> Allows customized variance-covariance matrices

- Key options are –w and –v
- Describe two alternative models for variances

Can build means model for association tests

- The –a option and –m options
- Using observed marker genotypes

Reads in IBD matrix generated by Merlin

Common Variance Components in QTDT

Value	Description
•	Non-shared Environment. Environmental effects that are
е	unique to each family member and measurement error.
a	Polygenic. These effects are a function of relatedness
g	between family members and may be due to polygenes.
	Additive Major Gene Effect. This represents the additive
а	effect of linkage to a major gene. The pi-hat component.
t	Twin Environment. This represents the environment shared
L	by twins, but not other types of relatives.
	Common Environment. This represents the environment
C	shared by all relatives.

The **–w** option specifies variances under the null. E.g. **–we** switch specifies that only environmental effects should be modeled. The **–v** switch specifies variances under the alternative. E.g. **–veg** models environmental and polygenic effects.

For other options and components, see documentation.

Testing Heritability With QTDT

Disable association models
 Use option -a-

Specify two models for variances
 Use options -we and -veg

> Typical command line: qtdt -d keavney.dat -p keavney.ped -a- -we -veg

Typical Output: Testing Heritability

Summary output appears on screen

The following models will be evaluated...

NULL MODEL Means = Mu Variances = Ve FULL MODEL Means = Mu Variances = Ve + Vg

Testing trait:

Allele	df(0) -	-LnLk(0)	df(V)	-LnLk(V)	ChiSq	р		
N/A	403	573.67	402	544.77	57.80	3e-14	(405	probands)

ACE

> In this case the trait is highly heritable ($p < 10^{-13}$)

> Additional information, including variance component estimates output to "regress.tbl"

Estimating IBD Coefficients

QTDT has very limited IBD engine More sophisticated estimates from Merlin are better

Relevant Merlin options

- Option --ibd generates merlin.ibd output file
- Option --markerNames to include marker labels

Use the following Merlin command line merlin --ibd --markerNames -d keavney.dat -p keavney.ped -m keavney.map

Testing Linkage With QTDT

Disable association models

 Use option -a

 Specify two models for variances

 Use options -we and -veg

 Provide an IBD file

 Use option -i ibdfile

> Typical command line: qtdt -d keavney.dat -p keavney.ped -i keavney.ibd -a- -weg -vega

Typical Output: Testing Linkage

The following models will be evaluated...

NULL MODEL Means = Mu Variances = Ve + Vg

FULL MODEL Means = Mu Variances = Ve + Vg + Va

Testing trait:

 Testing marker:
 T-5491C

 Allele
 df(0) -LnLk(0)
 df(V) -LnLk(V)
 ChiSq
 p

 All
 402
 544.77
 401
 528.22
 33.09
 9e-09
 (405 probands)

 Testing marker:
 A-5466C

 Allele
 df(0) -LnLk(0)
 df(V) -LnLk(V)
 ChiSq
 p

 Allele
 df(0) -LnLk(0)
 df(V) -LnLk(V)
 ChiSq
 p

 All
 402
 544.77
 401
 528.22
 33.09
 9e-09
 (405 probands)

ACE

(... output continues at each marker ...)

(... additional information in regress.tbl file ...)

Simple Association Model

Each copy of allele changes trait by a fixed amount
 Use covariate counting copies for allele of interest
 Results in estimate of additive genetic value
 Evidence for association when a ≠ 0

 $E(y_i) = \mu + a * [number of copies of mutantallele]$ $E(y_i) = \mu + \beta_X X_i$

X is the number of copies for allele of interest. β_x is the estimated effect of each copy (the additive genetic value).

Using QTDT to Test Association

Select association model

If population is homogenous, use option -at

Specify model for variances

- Use option –we if sample consists of unrelated individuals
- Use option -weg if sample consists of trios only
- Use option -wega under the null if families in general

Provide an IBD file

Typical command line:

qtdt -d keavney.dat -p keavney.ped -i merlin.ibd

-at -wega

Typical Output: Testing Association

The following models will be evaluated...

NULL MODEL Means = Mu Variances = Ve + Vg + Va

FULL MODEL Means = Mu + X Variances = Ve + Vg + Va

Testing trait:

ACE

Testing	marker:		<u>T-5491C</u>				
Allele	df(0) -I	lnLk(0)	df(X) -1	lnLk(X)	ChiSq	р	
1	382	503.12	381	471.97	62.29	3e-015	(386 probands)
2	382	503.12	381	471.97	62.29	3e-015	(386 probands)

Testing	marker:		A-54			
Allele	df(0) -LnLk(0	df(X)	-LnLk(X)	ChiSq	р	
1	380 503.	.1 37	9 471.07	64.09	1e-015	(384 probands)
2	380 503.	.1 37	9 471.07	64.09	1e-015	(384 probands)

(... output continues at each marker ...)

(... additional information in regress.tbl file ...)

Using QTDT to Test Association II

Select association model

• For the between-within model, use option -ao

Specify model for variances

- Use option -we if sample consists of unrelated individuals
- Use option -weg if sample consists of trios only
- Use option –wega under the null if families in general

Provide an IBD file

> Typical command line: qtdt -d keavney.dat -p keavney.ped -i merlin.ibd -ao -wega

Typical Output: Between-Within Association Model

ACE

The following models will be evaluated...

NULL MODEL Means = Mu + B Variances = Ve + Vg + Va

FULL MODEL Means = Mu + B + W Variances = Ve + Vg + Va

Testing trait:

 Testing marker:
 T-5491C

 Allele
 df(0) -LnLk(0)
 df(T) -LnLk(T)
 ChiSq
 p

 1
 381
 490.93
 380
 470.41
 41.04
 1e-010
 (180/386 probands)

 2
 381
 490.93
 380
 470.41
 41.04
 1e-010
 (180/386 probands)

<u>Testing</u>	marker:			A-54	66C		
Allele	df(0) -1	LnLk(0)	df(T) -1	lnLk(T)	ChiSq	р	
1	379	489.59	378	470.22	38.73	5e-010	(186/384 probands)
2	379	489.59	378	470.22	38.73	5e-010	(186/384 probands)

(... output continues at each marker ...)
(... additional information in regress.tbl file ...)

Can association explain linkage?

If linkage and association signals are present...
Using the previous two tests

Decide whether locus has been mapped
 Are there other associated alleles to be found?

Can the small region where LD extends ...

... account for the covariance between relatives who share the surrounding stretch of DNA?

Using QTDT to Test If Association Can Explain Linkage

Select an association model

For example, use option -at

Specify models for variances as in linkage test

- Use option -weg under the null
- Use option –vega under the alternative

Provide an IBD file

> Typical command line: qtdt -d keavney.dat -p keavney.ped -i merlin.ibd -at -weg -vega

Typical Output: Linkage after modeling association

The following models will be evaluated...

NULL MODEL Means = Mu + X Variances = Ve + Vg

FULL MODEL Means = Mu + X

Variances = Ve + Vg + Va

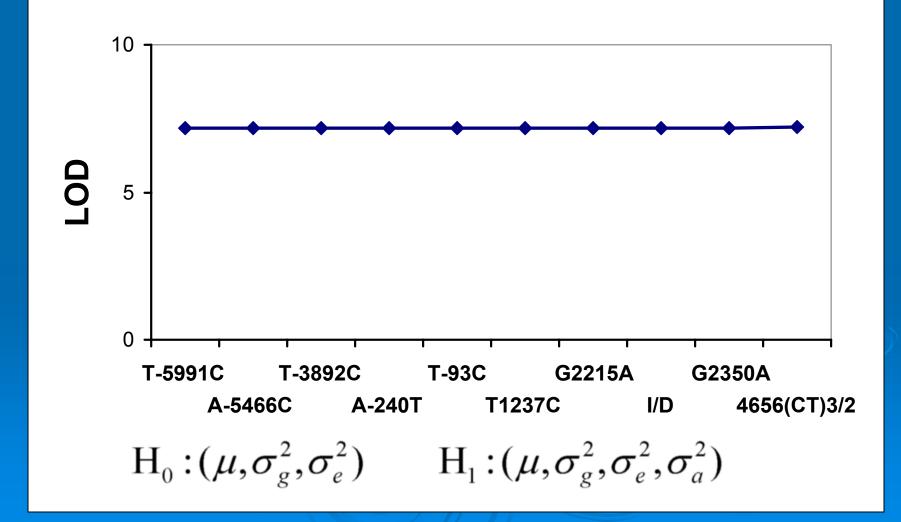
Testing trait:

ACE

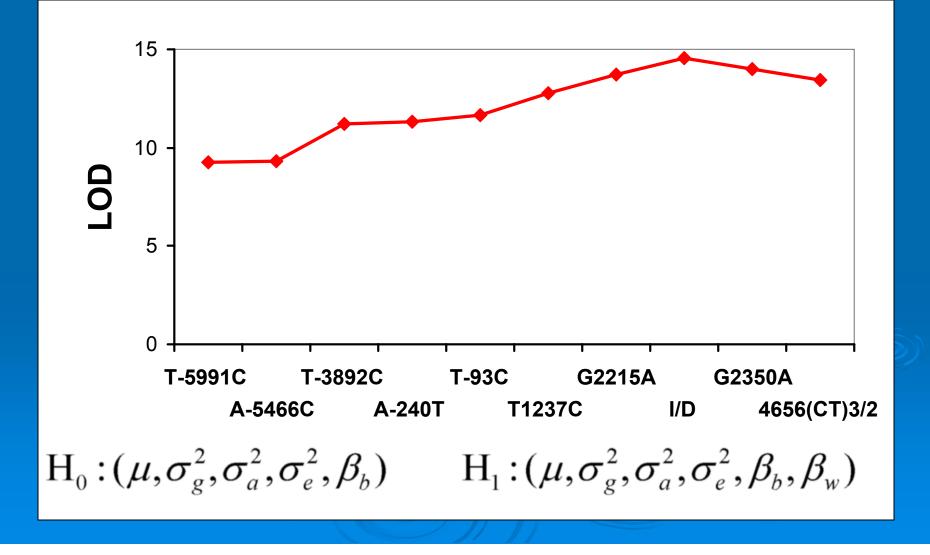
Testing	marker:			T-54			
Allele	df(0) -I	lnLk(0)	df(V) -1	lnLk(V)	ChiSq	р	
1	391	486.44	390	481.40	10.08	0.0015	(395 probands)
2	391	486.44	390	481.40	10.08	0.0015	(395 probands)

Testing	marker:		A-5466C				
Allele	df(0) -1	LnLk(0)	df(V) -1	LnLk(V)	ChiSq	р	
1	388	483.20	387	479.50	7.41	0.0065	(392 probands)
2	388	483.20	387	479.50	7.41	0.0065	(392 probands)

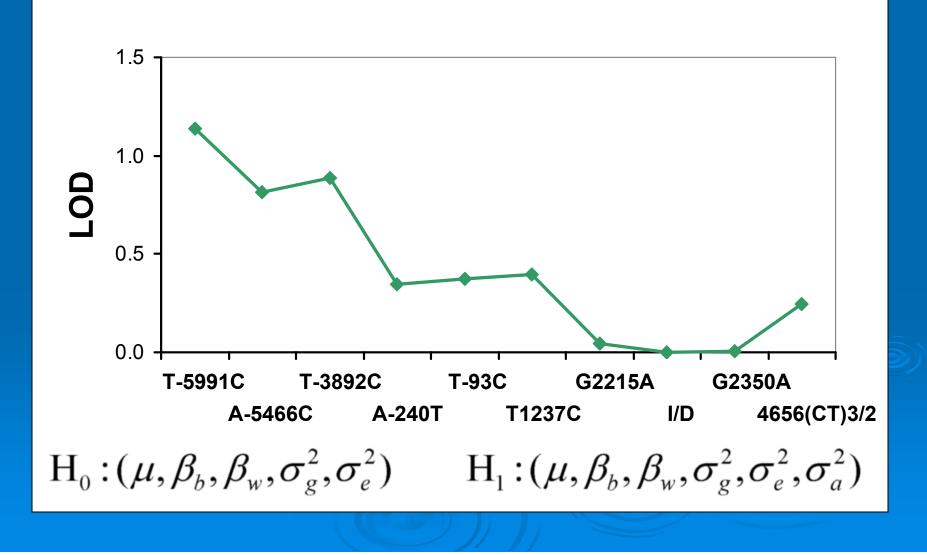
Evidence for Linkage



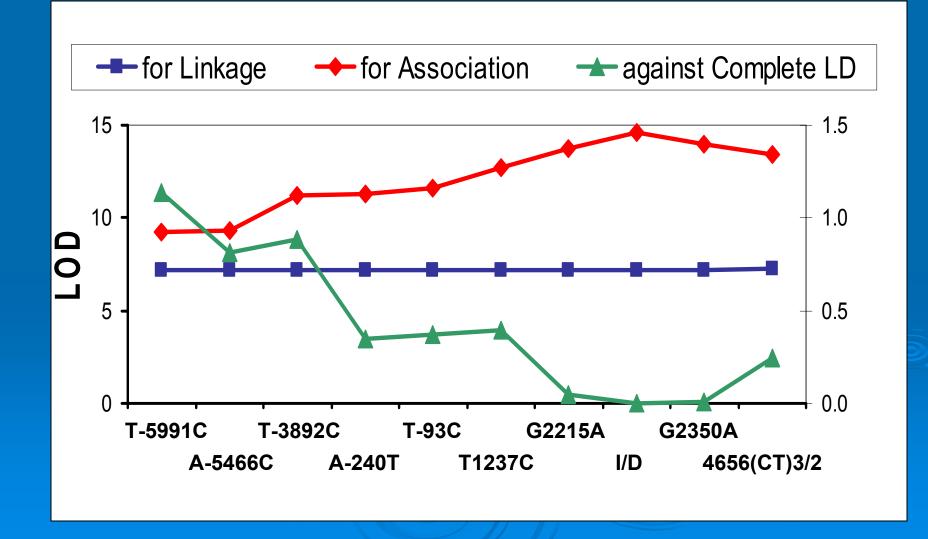
Evidence for Association



Evidence Against Complete LD



Drawing Conclusions



References

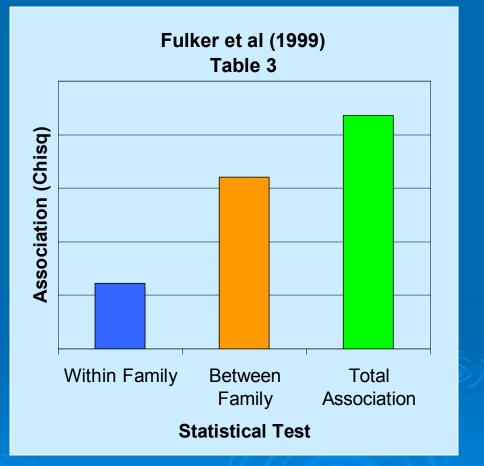
Fulker et al (1999) Am J Hum Genet 66:259-267 Neale et al (1999) Behav Genet 29: 233-244 Abecasis et al (2000) Am J Hum Genet 66:279-292 Sham et al (2000) Am J Hum Genet 66:1616-1630 Abecasis et al (2000) Eur J Hum Genet 8:545-551 Cardon and Abecasis (2000) Behav Genet 30:235-243

Family Based Association Studies Have Been Getting a Bad Rap...

- The most popular methods for association mapping in families rely on transmission disequilibrium
 TDT, sib-TDT, QTDT, PDT, ...
- Out of an abundance of caution, TDT-like tests use very conservative approach to guard against population structure
 - Reduced power on per genotype basis
- Given thousands of genotyped markers, better methods for guarding against stratification are available
 - Genomic Control (Devlin and Roeder, 1999)
 - Structured Association Mapping (Pritchard et al, 2000)
 - Avoid heavily stratified samples

Powerful Association Tests in Families

- Do not focus on within family association (e.g. transmission disequilibrium)
- Evaluate the effect of each genotype, adjusting for familial correlations (as in George and Elston 1987)
- Fulker et al (1999) showed that focusing on alleles transmitted from heterozygous parents discards 50-75% of information in a family sample



Transmission disequilibrium tests focus on within family component of association

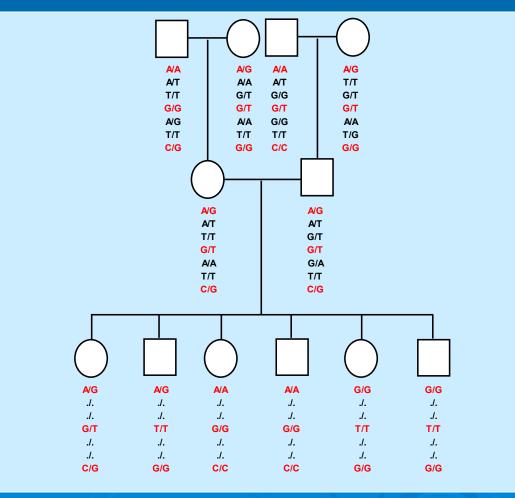
Incorporating Family Information in Genome Wide Studies

- Family members will share large segments of chromosomes
- If we genotype many related individuals, we will effectively be genotyping a few chromosomes many times

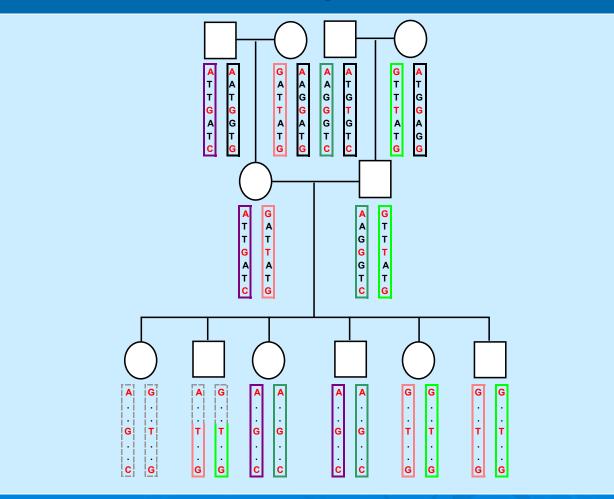
> In fact, we can:

- genotype a few markers on all individuals
- use high-density panel to genotype a few individuals
- infer shared segments and then estimate the missing genotypes

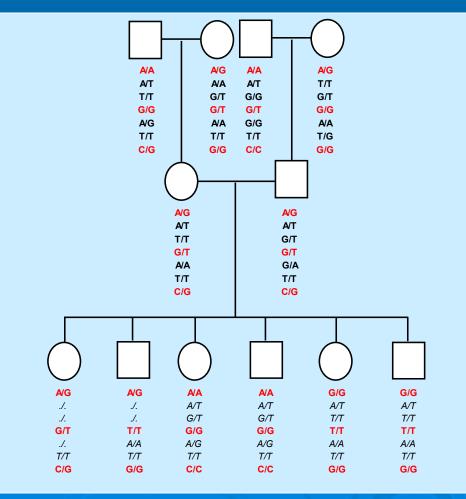
Genotype Inference Part 1 – Observed Genotype Data



Genotype Inference Part 2 – Inferring Allele Sharing



Genotype Inference Part 3 – Imputing Missing Genotypes



Our Approach

Consider full set of observed genotypes G

Evaluate pedigree likelihood L for each possible value of each missing genotype g_{ij}

Posterior probability for each missing genotype

$$P(g_{ij} = x | G) = \frac{L(G, g_{ij} = x)}{L(G)}$$

Implemented both using Elston-Stewart (1972) and Lander-Green (1987) algorithms

Standard Linear Model for Genetic Association

Model association using a model such as:

$$E(y_i) = \mu + \beta_g g + \beta_c c + \dots$$

- > y_i is the phenotype for individual *i*
- > g_i is the genotype for individual *i*
 - Simplest coding is to set g_i = number of copies of allele '1'
- > c_i is a covariate for individual *i*
 - Covariates could be estimated ancestry, environmental factors...
- β coefficients are estimated covariate, genotype effects
 Model is fitted in variance component framework

Model With Inferred Genotypes

> Replace genotype score g with its expected value:

$$E(y_i) = \mu + \beta_g \overline{g} + \beta_c c + \dots$$

> Where

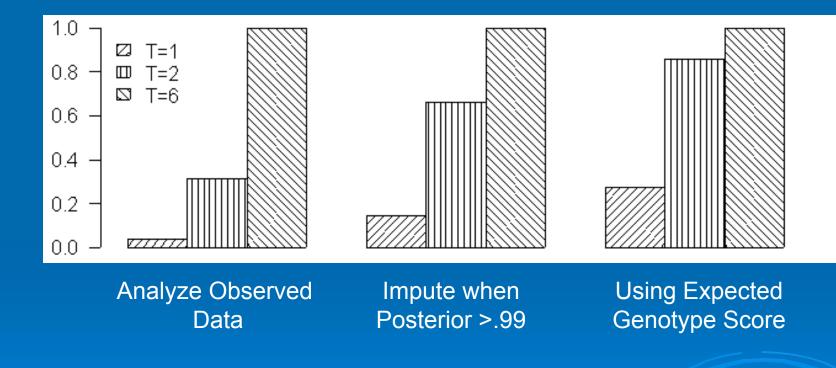
$$\overline{g}_i = 2P(g_i = 2 | G) + P(g_i = 1 | G)$$

Association test can then be implemented as a score test or as a likelihood ratio test

> Alternatives would be to

- (a) impute genotypes with large posterior probabilities; or
- (b) integrate joint distribution of unobserved genotypes in family

Power in Sibships of Size 6 Without Parental Genotype Data



T is the number of genotyped offspring. QTL explains 5% of variance, polygenes explain 35%, 250 sibships, $\alpha = 0.001$.

Application: Gene Expression Data

Cheung et al (2005) carried out a genome wide association with 27 expression levels as traits

Measured in grandparents and parents of CEPH pedigrees and took advantage of HapMap I genotypes

TSC genotypes also available for ~6000 SNPs in the offspring of each CEPH family

Results: Gene Expression Data

Using observed genotypes, the most significant association mapped in *cis* for 15 of 27 traits
 12 of these reach p < 5 * 10⁻⁸

Using inferred genotypes, the most significant association mapped in *cis* for 20 of 27 traits

- 15 of these reach $p < 5 * 10^{-8}$
- 1 trans linkage also reaches $p < 5 * 10^{-8}$

Sardinia

> 6,148 Sardinians from 4 towns in Ogliastra

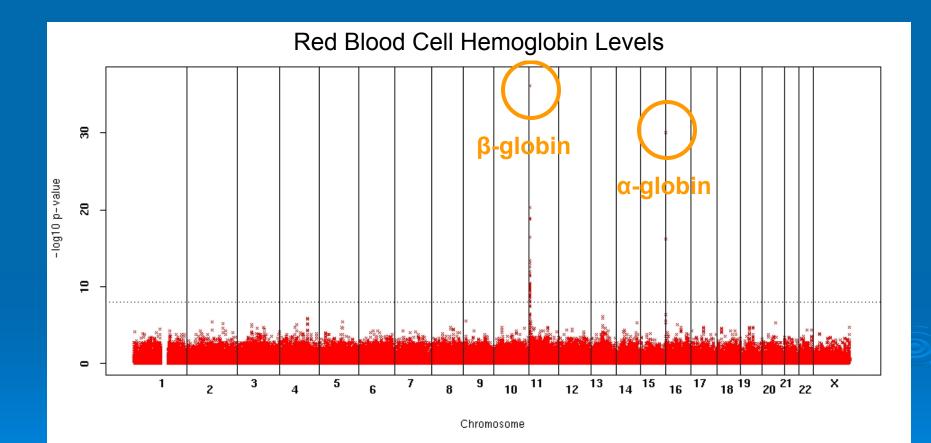
Measured 98 aging related quantitative traits

> Genotyping:

- Affymetrix 10K chip in 4,500 individuals
- Affymetrix 500K chip in 1,500 individuals

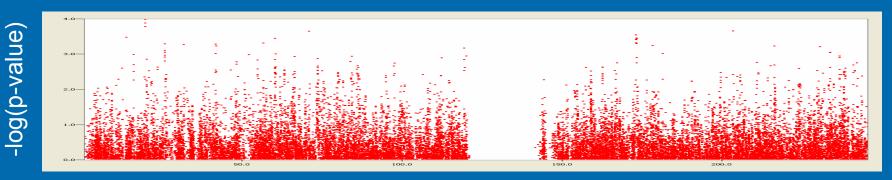
Large pedigrees, computationally challenging
 Preliminary results

Preliminary Results: Sardinia (after ~900 500K arrays)



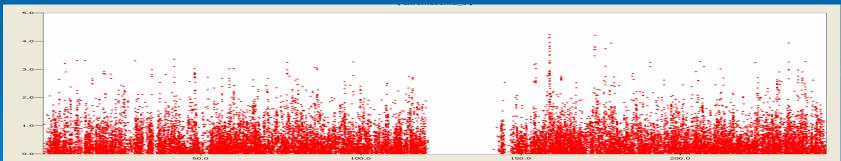
Preliminary Results from Sardinia QT interval, Chromosome 1

Before imputation



After imputation





Position (in Mb) Along Chromosome 1

Acknowledgements

> Weimin Chen, Serena Sanna

 Sardinia Investigators, led by:
 David Schlessinger, Manuela Uda, Antonio Cao, Edward Lakatta, Paul Costa

Gene Expression Data:
 Vivian Cheung, Josh Burdick

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