

REPEATED MEASURES ANOVA

Repeated measures ANOVA (RM) is a specific type of MANOVA. When the within group covariance matrix has a special form, then the RM analysis usually gives more powerful hypothesis tests than does MANOVA. Mathematically, the within group covariance matrix is assumed to be a type H matrix (SAS terminology) or to meet Huynh-Feldt conditions. These mathematical conditions are given in the appendix. Most software for RM prints out both the MANOVA results and the RM results along with a test of RM assumption about the within group covariance matrix. Consequently, if the assumption is violated, one can interpret the MANOVA results. In practice, the MANOVA and RM results are usually similar.

There are certain stock situations when RM is used. The first occurs when the dependent variables all measure the same construct. Examples include a time series design of growth curves of an organism or the analysis of number of errors in discrete time blocks in an experimental condition. A second use of RM occurs when all the dependent variables are all measured on the same scale (e.g., the DVs are all Likert scale responses). For example, outcome from therapy might be measured on Likert scales reflecting different types of outcome (e.g., symptom amelioration, increase in social functioning, etc.). A third situation is for internal consistency analysis of a set of items or scales that purport to measure the same construct.. Internal consistency analysis consists in fitting a linear model to a set of items that are hypothesized to measure the same construct. For example, suppose that you have written ten items that you think measure the construct of empathy. Internal consistency analysis will provide measures of the extent to which these items "hang together" statistically and measure a single construct. (NOTE WELL: as in most stats, good internal consistency is just an index; *you* must be the judge of whether the single construct is empathy or something else.) If you are engaged in this type of scale construction, you should also use the SPSS subroutine RELIABILITY or SAS PROC CORR with the ALPHA option.

The MANOVA output from a repeated measures analysis is similar to output from traditional MANOVA procedures. The RM output is usually expressed as a "univariate" analysis, despite the fact that there is more than one dependent variable. The univariate RM has a jargon all its own which we will now examine by looking at a specific example.

An example of an RM design

Consider a simple example in which 60 students studying a novel foreign language are randomly assigned to two conditions: a control condition in which the foreign language is taught in the traditional manner and a experimental condition in which the language is taught in an

“immersed” manner where the instructor speaks only in the foreign language. Over the course of a semester, five tests of language mastery are given. The structure of the data is given in Table 1.

Table 1. Structure of the data for a repeated measures analysis of language instruction.

Student	Group	Test 1	Test 2`	Test 3	Test 4	Test 5
Abernathy	Control	17	22	26	28	31
.
Zelda	Control	18	24	25	30	29
Anasthasia	Experimental	16	23	28	29	34
.
Zepherinus	Experimental	23	25	29	38	47

The purpose of such an experiment is to examine which of the two instructional techniques is better. One very simple way of doing this is to create a new variable that is the sum of the 5 test scores and perform a *t*-test. The SAS code would be

```
DATA rmex1;
  INFILE 'c:\sas\p7291dir\repeated.simple.dat';
  LENGTH group $12.;
  INPUT subjnum group test1-test5;
  testtot = sum(of test1-test5);
RUN;

PROC TTEST;
  CLASS group;
  VAR testtot;
RUN;
```

The output from this procedure would be:

TTEST PROCEDURE
Variable: TESTTOT

GROUP	N	Mean	Std Dev	Std Error
Control	30	154.4333333	27.96510551	5.10570637
Experimental	30	170.5000000	32.91237059	6.00894926
Variances	T	DF	Prob> T	
Unequal	-2.0376	56.5	0.0463	
Equal	-2.0376	58.0	0.0462	

For H0: Variances are equal, F' = 1.39 DF = (29,29) Prob>F' = 0.3855

The mean total test score for the experimental group (170.5) is greater than the mean total test score for controls (154.4). The difference is significant ($t = -2.04$, $df = 58$, $p < .05$) so we should conclude that the experimental language instruction is overall superior to the traditional language instruction.

There is nothing the matter with this analysis. It gives an answer to the major question posed by the research design and suggests that in the future, the experimental method should be adopted for foreign language instruction.

But the expense of time in generating the design of the experiment and collecting the data merit much more than this simple analysis. One very interesting question to ask is whether the means for the two groups change over time. Perhaps the experimental instruction is good initially but poor at later stages of foreign language acquisition. Or maybe the two techniques start out equally but diverge as the semester goes on. A simple way to answer these questions is to perform separate t -tests for each of the five tests. The code here would be:

```
PROC TTEST DATA=r mex1;
  CLASS GROUP;
  VAR test1-test5;
RUN;
```

A summary of these results is given in Table 2.

Table 2. Group means and t -test results for the 5 tests.				
Means:				
Test	Control	Experimental	t	$p <$
1	20.17	18.80	.70	.49
2	27.8	29.1	-.59	.56
3	30.7	36.1	-2.66	.01
4	36.83	40.43	-1.90	.07
5	38.93	46.07	-3.35	.002

At this stage a plot of the means is helpful.

Both the plot of the means and the t -test results suggest that the two groups start out

fairly similar at times 1 and times 2, diverge significantly at time 3, are almost significantly different at time 4, and diverge again at time 5. This analysis gives more insight, but leads to its own set of problems. We have performed 5 different significance tests. If these tests were independent—and they are clearly not independent because they were performed on the same set of individuals—then we should adjust the α level by the Bonferroni formula

$$\alpha_{\text{adjusted}} = 1 - .95^{\frac{1}{\text{number of tests}}} = 1 - .95^{.2} = .01.$$

Using this criterion, we would conclude that the differences in test 3 are barely significant, those in test 4 are not significant, while the means for the last test are indeed different. Perhaps the major reason why the two groups differ is only on the last exam in the course.

A repeated measures example can help to clarify the situation. The advantage to RM is that it will control for the correlations among the tests and come up with an overall test for each of the hypotheses given above. The RM design divides ANOVA factors into two types: *between subjects factors* (or effects) and *within subject factors* (or effects). If you think of the raw data matrix, you should have little trouble distinguishing the two. A single between subjects factor has one and only one value per observation. Thus, *group* is a between subjects factor because each observation in the data matrix has only one value for this variable (*Control* or *Experimental*).

Within subjects factors have more than one value per observation. Thus, *time of test* is a within subject factor because each subject has five difference values--Test1 through Test5. Another way to look at the distinction is that between subject factors are all the independent variables. Within subject factors involve the dependent variables. All interaction terms that involve a within subject factor are included in within subject effects. Thus, interactions of *time of test* with *group* is a within subject effect. Only interactions that include *only* between subject factors are included in between subjects effects.

The effects in a RM ANOVA are the same as those in any other ANOVA. In the present example, there would be a main effect for *group*, a main effect for *time* and an interaction between *group* and *time*. RM differs only in the mathematics used to compute these effects.

At the expense of putting the cart before the horse, the SAS commands to perform the repeated measures for this example are:

```
TITLE Repeated Measures Example 1;
PROC GLM DATA=rhex1;
  CLASS group;
  MODEL test1-test5 = group;
  MEANS group;
  REPEATED time 5 polynomial / PRINTM PRINTE SUMMARY;
RUN;
```

As in an ANOVA or MANOVA, the CLASS statement specifies the classification variable which is *group* in this case. The MODEL statement lists the dependent variables (on the left hand side of the equals sign) and the independent variables (on the right hand side). The MEANS statement asks that the sample size, means, and standard deviations be output for the two groups.

The novel statement in this example is the REPEATED statement. Later, this statement will be discussed in detail. The current REPEATED statement gives a name for the repeated measures or within subjects factor—*time* in this case—and the number of levels of that factor—5 in this example because the test was given over 5 time periods. The word *polynomial* instructs SAS to perform a polynomial transform of the 5 dependent variables. Essentially this creates 4 “new” variables from the original 5 dependent variables. (Any transformation of k dependent variables will result in $k - 1$ new transformed variables.) The PRINTM option prints the transformation matrix, the PRINTE option prints the error correlation matrix and some other important output, and the SUMMARY option prints ANOVA results for each of the four transformed variables.

Usually it is the transformation of the dependent variables that gives the RM analysis additional insight into the data. Transformations will be discussed at length later. Here we just note that the *polynomial* transformation literally creates 4 new variables from the 5 original dependent variables. The first of the new variables is the linear effect of time; it tests whether the means of the language mastery tests increase or decrease over time. The second new variable is the quadratic effect of time. This new variable tests whether the means have a single “bend” to them over time. The third new variable is the cubic effect over time; this tests for two “bends” in the plot of means over time. Finally the fourth new variable is the quartic effect over time, and it tests for three bends in the means over time.

The first few pages of output from this procedure give the results from the univariate ANOVAs for test1 through test5. Because there are only two groups, the F statistics for these analysis are equal to the square of the t statistics in Table 2 and the p values for the ANOVAs will be the same as those in Table 2. Hence these results are not presented. The rest of the output begins with the MEANS statement.

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 General Linear Models Procedure

Level of		-----TEST1-----		-----TEST2-----	
GROUP	N	Mean	SD	Mean	SD
Control	30	20.1666667	7.65679024	27.8000000	7.76996871
Experimental	30	18.8000000	7.46208809	29.1000000	9.12499410

Level of		-----TEST3-----		-----TEST4-----	
GROUP	N	Mean	SD	Mean	SD
Control	30	30.7000000	6.88902175	36.8333333	7.06659618
Experimental	30	36.1000000	8.70334854	40.4333333	7.57347155

Level of		-----TEST5-----	
GROUP	N	Mean	SD
Control	30	38.9333333	8.65401375
Experimental	30	46.0666667	7.80332968

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 General Linear Models Procedure
 Repeated Measures Analysis of Variance
 Repeated Measures Level Information

The following section of output shows the design of the repeated measure factor. It is quite simple in this case. You should always check this to make certain that design specified on the REPEATED statement is correct.

Dependent Variable	TEST1	TEST2	TEST3	TEST4	TEST5
Level of TIME	1	2	3	4	5

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 Repeated Measures Analysis of Variance

These are the partial correlations for the dependent variables controlling for all the independent variables in the model. They answer the question, "To what extent is Test1 correlated with Test2 within each of the two groups?" Usually, these correlations should be will be significant.

Partial Correlation Coefficients from the Error SS&CP Matrix / Prob > |r|

DF = 58	TEST1	TEST2	TEST3	TEST4	TEST5
TEST1	1.000000 0.0001	0.451725 0.0003	0.417572 0.0010	0.510155 0.0001	0.477928 0.0001
TEST2	0.451725 0.0003	1.000000 0.0001	0.445295 0.0004	0.599058 0.0001	0.493430 0.0001
TEST3	0.417572 0.0010	0.445295 0.0004	1.000000 0.0001	0.650573 0.0001	0.465271 0.0002
TEST4	0.510155 0.0001	0.599058 0.0001	0.650573 0.0001	1.000000 0.0001	0.493323 0.0001
TEST5	0.477928 0.0001	0.493430 0.0001	0.465271 0.0002	0.493323 0.0001	1.000000 0.0001

Below is the transformation matrix. It is printed here because the PRINTM option was specified in the REPEATED statement. Because we specified a POLYNOMIAL transformation, this matrix gives coefficients for what are called *orthogonal polynomials*. They are analogous but not identical to contrast codes for independent variables. The first new variable, TIME.1, gives the linear effect over time, the second, TIME.2 is the quadratic effect, etc.

TIME.N represents the nth degree polynomial contrast for TIME

M Matrix Describing Transformed Variables

	TEST1	TEST2	TEST3	TEST4	TEST5
TIME.1	-.6324555320	-.3162277660	0.0000000000	0.3162277660	0.6324555320
TIME.2	0.5345224838	-.2672612419	-.5345224838	-.2672612419	0.5345224838
TIME.3	-.3162277660	0.6324555320	-.0000000000	-.6324555320	0.3162277660
TIME.4	0.1195228609	-.4780914437	0.7171371656	-.4780914437	0.1195228609

E = Error SS&CP Matrix

TIME.N represents the nth degree polynomial contrast for TIME

	TIME.1	TIME.2	TIME.3	TIME.4
TIME.1	1723.483333	-2.521377	236.550000	313.306091
TIME.2	-2.521377	2187.726190	98.502728	35.622692
TIME.3	236.550000	98.502728	1657.950000	-468.112779
TIME.4	313.306091	35.622692	-468.112779	1715.000476


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General Linear Models Procedure
Repeated Measures Analysis of Variance

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The within group partial correlation matrix, but this time for the transformed variables. If the transformation resulted in orthogonal variables, then the test for Sphericity applied to the orthogonal components should be non significant. If the test IS significant, then interpret the MANOVA results instead of the repeated measured results.

Partial Correlation Coefficients from the Error SS&CP Matrix
of the Variables Defined by the Specified Transformation / Prob > |r|

DF = 58	TIME.1	TIME.2	TIME.3	TIME.4
TIME.1	1.000000	-0.001298	0.139937	0.182236
	0.0001	0.9922	0.2905	0.1671
TIME.2	-0.001298	1.000000	0.051721	0.018391
	0.9922	0.0001	0.6972	0.8900
TIME.3	0.139937	0.051721	1.000000	-0.277608
	0.2905	0.6972	0.0001	0.0333
TIME.4	0.182236	0.018391	-0.277608	1.000000
	0.1671	0.8900	0.0333	0.0001

Test for Sphericity: Mauchly's Criterion = 0.8310494
Chisquare Approximation = 10.440811 with 9 df Prob > Chisquare = 0.3160

Applied to Orthogonal Components:
Test for Sphericity: Mauchly's Criterion = 0.8310494
Chisquare Approximation = 10.440811 with 9 df Prob > Chisquare = 0.3160

Manova Test Criteria and Exact F Statistics for
the Hypothesis of no TIME Effect
H = Type III SS&CP Matrix for TIME E = Error SS&CP Matrix

S=1 M=1 N=26.5

Statistic	Value	F	Num DF	Den DF	Pr > F
Wilks' Lambda	0.07669737	165.5260	4	55	0.0001
Pillai's Trace	0.92330263	165.5260	4	55	0.0001
Hotelling-Lawley Trace	12.03825630	165.5260	4	55	0.0001
Roy's Greatest Root	12.03825630	165.5260	4	55	0.0001

```

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Manova Test Criteria and Exact F Statistics for
the Hypothesis of no TIME*GROUP Effect
H = Type III SS&CP Matrix for TIME*GROUP    E = Error SS&CP Matrix

S=1      M=1      N=26.5

Statistic              Value              F          Num DF      Den DF    Pr > F
Wilks' Lambda          0.74044314          4.8200           4          55    0.0021
Pillai's Trace          0.25955686          4.8200           4          55    0.0021
Hotelling-Lawley Trace  0.35054260          4.8200           4          55    0.0021
Roy's Greatest Root     0.35054260          4.8200           4          55    0.0021
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General Linear Models Procedure
Repeated Measures Analysis of Variance
Tests of Hypotheses for Between Subjects Effects

Source              DF      Type III SS      Mean Square      F Value      Pr > F
GROUP                1          774.4133          774.4133          4.15          0.0462
Error                58        10818.5733          186.5271

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General Linear Models Procedure
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Univariate Tests of Hypotheses for Within Subject Effects

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Source: TIME

DF	Type III SS	Mean Square	F Value	Pr > F	Adj G - G	Pr > F H - F
4	19455.82000000	4863.95500000	154.92	0.0001	0.0001	0.0001

Source: TIME*GROUP

DF	Type III SS	Mean Square	F Value	Pr > F	Adj G - G	Pr > F H - F
4	674.02000000	168.50500000	5.37	0.0004	0.0005	0.0004

Source: Error(TIME)

DF	Type III SS	Mean Square
232	7284.16000000	31.39724138

Greenhouse-Geisser Epsilon = 0.9332

Huynh-Feldt Epsilon = 1.0225

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Analysis of Variance of Contrast Variables

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TIME.N represents the nth degree polynomial contrast for TIME

Contrast Variable: TIME.1

Source	DF	Type III SS	Mean Square	F Value	Pr > F
MEAN	1	18961.88167	18961.88167	638.12	0.0001
GROUP	1	558.73500	558.73500	18.80	0.0001
Error	58	1723.48333	29.71523		

Contrast Variable: TIME.2

Source	DF	Type III SS	Mean Square	F Value	Pr > F
MEAN	1	421.4583333	421.4583333	11.17	0.0015
GROUP	1	18.6011905	18.6011905	0.49	0.4853
Error	58	2187.7261905	37.7194171		

Contrast Variable: TIME.3

Source	DF	Type III SS	Mean Square	F Value	Pr > F
MEAN	1	42.13500000	42.13500000	1.47	0.2296
GROUP	1	22.81500000	22.81500000	0.80	0.3753
Error	58	1657.9500000	28.58534483		

Contrast Variable: TIME.4

Source	DF	Type III SS	Mean Square	F Value	Pr > F
MEAN	1	30.34500000	30.34500000	1.03	0.3152
GROUP	1	73.86880952	73.86880952	2.50	0.1194
Error	58	1715.0004762	29.56897373		

The RM design divides ANOVA factors into two types: *between subjects* factors (or effects) and *within subject* factors (or effects). If you think of the above type of data matrix, you should have little trouble distinguishing the two. Between subjects factors have one and only one value per observation. Thus, Mode, Type, and Age are between subjects factors. Within subjects factors have more than one value per observation. Thus, Test is a within subject factor because each subject has five difference values--Test1 through Test5. Another way to look at the distinction is that between subject factors are all the independent variables. Within subject factors are the dependent variables. All interaction terms that involve a within subject factor are included in within subject effects. Thus, interactions of test with mode, test with type, and test with age would be considered a within subject effect. Only interactions that include *only* between subject factors are included in between subjects effects. Thus, the interaction of mode with type or mode with age are between subject effects.

The between subjects effects answers the following question: why does one subject's average score over the five tests differ from another subject's average score? In the above design, there are several possible reasons--the effect of Mode, the effect of Type, the effect of Age, and any interactions among these three. The within subjects effects ask why the five tests might differ for any single subject. There are several reasons. First, the tests might differ in difficulty--a Test effect. One Mode of instruction might make it easier to master some parts of the language earlier than others--a Test by Mode interaction. Some Types might affect early versus late mastery--a Test by Type interaction. And learning curves might depend upon age--an Age by Test interaction. One could even construct higher order interactions, say a Test by Mode by Type by Age interaction. All of these interactions are included in the within subjects effects because they all include an interaction with Test, a within subjects factor.

In order to test for univariate within subjects effects using RM, it is necessary to transform the dependent variables. We have already seen one type of transformation of dependent variables in profile analysis in MANOVA. There are several other types of transformations useful for RM. They are discussed below in the section on transformations.

Transformations: Why Do Them?

The primary reason for transforming the dependent variables in RM is to generate variables that are more informative than the original variables. Consider the RM example. The dependent variables are tests of language mastery. Scores on a mastery test should be low early in

the course and increase over the course. At some point, we might expect them to asymptote, depending upon the difficulty of the mastery test. We can then reformulate the research questions into the following set of questions: Do some types of instruction increase mastery at a faster rate than other types? Does computer based instruction increase mastery at a faster rate than classroom instruction? To answer these questions, we want to compare the independent variables on the *linear* increase over time in mastery test scores. If the mastery test is constructed so that scores asymptote at some time point during the course, we could ask the following questions: Do some types of instructions asymptote faster than other types? Or, does computer based instruction asymptote faster than classroom instruction? To answer these questions, we want to test the differences in the *quadratic* effect over time for the independent variables. In short, we want to transform the test scores so that the first transformed variable is the linear effect over time and the second variable is the quadratic effect over time. We can then do a MANOVA or RM on the transformed variables.

There are also statistical reasons for transformations¹. We can see the statistical philosophy behind transformations by recalling the major reason for using MANOVA instead of interpreting a series of univariate analyses. We use MANOVA because dependent variables are correlated. If we get obsessed about the whole business and want to interpret a series of univariate analyses, then the best way of doing that is to transform the dependent variables in such a way that they are uncorrelated. This is one good statistical reason for transforming the variables in RM.

We also want the variances of the transformed variables to be homogeneous. The reason for this is that we would like to have one single overall F for the dependent variables to determine whether or not there are significant effects. We cannot get an exact F statistic unless the variances of the dependent variables are homogeneous. The situation is analogous to a simple univariate F

¹ The real justification for the transformation is to test whether the covariance matrix has a particular form that will make the F statistic exact. An identity matrix has this form and a matrix with homogeneous variances (all diagonal elements equal) and homogeneous covariances (all off diagonal elements equal) also has this form. But there are other matrices that will also permit repeated measures. One characteristic common to all these matrices is that an orthonormal transformation of the dependent variables gives an expected correlation matrix that is an identity matrix. The interested reader is referred to Huynh, H. and Feldt, L.S. (1970), Conditions under which mean square ratios in repeated measurements designs have exact F -distributions. *Journal of the American Statistical Association*, 65, 1582-1589.

test for, say, three groups. In order for the F test to be valid the variances of the four groups must be homogeneous. In RM, the variances of the dependent variables must be homogeneous. One way to achieve homogeneity of variances is to transform the dependent variables so that their variances are unity.

Thus, two ideal statistical requirements of transformation are that: (1) the variances are unity, and (2) the transformed variables are uncorrelated. If we can achieve this, we have an *orthonormalized* transformation. Some types of transformations will guarantee that the dependent variables will be orthonormalized--e.g., rescaled principal components. However, the transformations most often used in RM designs do not always guarantee that the transformed dependent variables will in fact be orthonormalized. Hence, we must always test whether the transformation has worked. In order to do this, we compare the transformed covariance matrix to an identity matrix using a test called a *sphericity* test. There are several types of sphericity tests, but all are estimates of a likelihood ratio c^2 . [The tests differ in how they adjust the c^2 for the fact that in small samples, the likelihood ratio c^2 is not asymptotically valid.] Both SAS and SPSSx will spit out sphericity tests upon request. If the test is significant, then the transformation has not worked. If the test is not significant, then most researchers assume the transformation has worked and will interpret the results of the univariate RM. Most sphericity tests are regarded as a "sensitive" test. That is, they will often give "statistical significance" even though there is little substantive difference between the correlation matrix for the transformed variables and an identity matrix.

If the transformation gives a significant c^2 , there are several options to choose from:

- (1) ignore it because the matrix is pretty close to an identity matrix. (Not recommended because picky journal editors and referees often get upset over this, even though there might not be anything wrong with it.)
- (2) use the MANOVA results to interpret the within subjects effects.
- (3) use the Greenhouse - Geisser correction or the Huynh - Feldt correction to the univariate F statistics. Both of these are adjustments to the degrees of freedom for an F statistic to take into account the fact that the covariance matrix does not meet the strict requirements of RM. SAS prints these two corrections and their associated F statistics under the respective labels 'G - G' and 'H - F'. SAS also prints out the Greenhouse - Geisser and the Huynh - Feldt . Both of these

statistics index the extent to which the transformed matrix meets the requirements of the RM design. In the ideal case where the requirements are exactly met to the n th decimal place, then ϵ should equal 1.0. The value of ϵ gets small and approaches 0 as the requirements are more and more violated.

Transformations: How to Do Them

Both SAS and SPSSx recognize that constructing transformation matrices is a real pain in the gluteus to the max. Thus, they do it for you. As a user of a RM design, your major obligation is to choose the transformation *that makes the most sense for your data*. SAS will automatically perform the following transformations for you:

CONTRAST. A CONTRAST transformation compares each level of the repeated measures with the first level. It is useful when the first level represents a control or baseline level of response and you want to compare the subsequent levels to the baseline. For our example, the first transformed variable will be the (Test1 - Test2), the second transformed variable will be (Test1 - Test3), the third (Test1 - Test4) and the last (Test1 - Test5). CONTRAST is the default transformation in SAS--the one you get if you do not specify a transformation.

MEAN. A MEAN transformation compares a level with the mean of all the other levels. It is mostly useful if you haven't the vaguest idea of how to transform the repeated measures variables. For the example, the first transformed variable will be (Test1 - mean of [Test2 + Test3 + Test4 + Test5]), the second variable will be (Test2 - mean of [Test1 + Test3 + Test4 + Test5]), etc. Note that there is always one less transformation than the number of variables. Hence, if you use a MEAN transform in SAS, you will not get the last level contrasted with the mean of the other levels. If you have a burning passion to do this, see the PROC GLM documentation in the SAS manual.

PROFILE. A PROFILE transformation compares a level against the next level. It is sometimes useful in testing responses that are not expected to increase or decrease regularly over time. For the example, the first transformed variable is (Test1 - Test2), the second is (Test2 - Test3), the third is (Test3 - Test4), etc.

HELMERT. A HELMERT transform compares a level to the mean of all subsequent levels. This is a very useful transformation when one wants to pinpoint when a response changes over time. For our example, the first transformed variable would be (Test1 - mean of [Test2 + Test3 + Test4 + Test5]), the second would be (Test2 - mean of [Test3 + Test4 + Test5]), the third would be (Test3 - mean of [Test4 + Test5]), and the last would simply be (Test4 - Test5). If the univariate *F* statistics were significant for the first and second transformed variables but not significant for the third and fourth, then we would conclude that language mastery was achieved

by the time of the third test.

POLYNOMIAL. A POLYNOMIAL transform fits orthogonal polynomials. Like a Helmert transform, this is useful to pinpoint changes in response over time. It is also useful when the repeated measures are ordered values of a quantity, say the dose of a drug. The first transformed variable represents the linear effect over time (or dose). The second transformed variable denotes the quadratic effect, the third the cubic effect, etc. If you are familiar enough with polynomials to interpret the observed means in light of linear, quadratic, cubic, etc. effects, this is an exceptionally useful transformation. Often, one can predict beforehand the order of the polynomial but not the exact time period where the response might be maximized (or minimized).

SPSSx will also transform repeated measures variables using the CONTRAST subcommand to the MANOVA procedure. Note, however, that terminology and procedures differ greatly between SAS and SPSSx. Be particularly careful of the CONTRAST statement. In SAS, the CONTRAST *statement* refers to a transformation of only the independent variables. The CONTRAST *option* on the REPEATED statement allows for a specific type of transformation for repeated measures variables. SPSSx views a CONTRAST as a transformation of either independent variables or dependent variables, depending upon the context. To make the issue more confusing, SPSSx has another subcommand, TRANSFORM, that applies only to the dependent variables. Also, both packages will do a "profile" transformation, but actually do different transformations. *You should always consult the appropriate manual before ever transforming the repeated measures variables.*

Interpreting Repeated Measures

To interpret repeated measures results, it is helpful to create a table of the effects. Suppose that the model we fit for our example was based on the following SAS statements:

```
PROC GLM DATA=INSTRUC;
  CLASS MODE TYPE;
  MODEL TEST1--TEST5 = MODE TYPE MODE*TYPE AGE;
```

Then we can construct a Table that looks like this:

<u><i>p</i> values from:</u>				
<u>ANOVA Effects</u>	<u>MANOVA F</u>	<u><i>F</i></u>	<u>GG <i>F</i></u>	<u>HF <i>F</i></u>
Between Subjects:				
MODE				
TYPE				
MODE*TYPE				
AGE				
Within Subjects:				
TIME				
MODE*TIME				
TYPE*TIME				
MODE*TYPE*TIME				
AGE*TIME				
ERROR				

Here the variable TIME is used to denote the set of the five tests.² We now want to fill in the *p* values from the various *F* statistics. The between subjects effects are equivalent to the effects from an ANOVA using an individual's average score (or total score) on the five tests as the dependent variable. Consequently, all the *F* statistics and their associated *p* values will be the

² Belatedly, I note that I switched terms here. The TIME effect in this analysis is the same as the TEST effect referred to in the previous sections.

same for the between subjects effects. For the within subjects effects, however, the F statistics and associated p values will generally differ.

The within subjects effects are interpreted as if they were ANOVA factors. The TIME effect within subjects tells us whether the means for the five tests are equal. The MODE*TIME within subject effect has the same interpretation as a MANOVA using MODE as the independent variable and the five tests (TIME) as the dependent variables. That is, are the (5 by 1) vector of means on the five tests for the Classroom condition and the (5 by 1) vector of means on the five tests for the Computer condition sampled from the same distribution? The TYPE*TIME within subject factor tests whether the three (5 by 1) vectors of means for the Empirical, Programmed, and Didactic condition are pulled from the same hat. The MODE*TYPE*TIME effect tests whether the six (5 by 1) vectors of means for the two Modes and three Types are different from those predicted on the basis of knowing a Mode effect and a Time effect. Finally, the AGE*TIME effect tests whether the five regression coefficients for AGE are the same for all five tests.

We can now fill in the table from the computer output. Here, I have chosen to put in the p value associated with Wilk's l for the multivariate F .

		<u>p values from:</u>		
<u>ANOVA Effects</u>	<u>MANOVA F</u>	<u>F</u>	<u>GG F</u>	<u>HF F</u>
Between Subjects:				
MODE		.31		
TYPE		.10		
MODE*TYPE		.80		
AGE		.0001		
Within Subjects:				
TIME	.03	.02	.02	.02
MODE*TIME	.56	.67	.66	.67
TYPE*TIME	.0001	.0001	.0001	.0001
MODE*TYPE*TIME	.66	.63	.62	.63
AGE*TIME	.99	.99	.99	.99
ERROR				

To interpret the table, we first need to examine the test of sphericity. For this problem, $c^2 = 12.0$ ($df = 9$, $p = .21$). This suggests that the correlation matrix among the variables has the appropriate form that permits the more powerful F test to be used. Were the c^2 significant, then we would ignore the column for the F and interpret the columns for the MANOVA, the Greenhouse-Geiser correction, and the Huynh-Feldt correction.

As it stands, all four test statistics yield the same result. If the four columns differed, then choose the column for F when the sphericity test is passed. If the test is not passed, then you must make a decision based on the other three columns, and there is no established criteria for choosing among the three. The G-G correction is the most conservative, especially in small samples. That is, you are less likely to reject a null hypothesis with this test than with the MANOVA or the H-F correction.

Now let's interpret the substance of these results. For the total score over the five tests (the between subjects factors), there is no effect for Mode, Type, or their interaction. This means that average mastery levels over the course of the semester does not depend upon the way in which students were instructed. Age, however, is significant, so average mastery is predicted by student's age.

For within subject factors, there is a significant Time effect. Thus, the means for the five tests differ, the means being based on everyone in the sample. There is no effect for Mode, suggesting that over time, mastery does not depend upon computer or classroom teaching. There is an effect for Type. This means that over time, the Type of instruction is producing mastery at different rates. There is no effect for the interaction of Mode and Test nor for age.

To examine the effect, it is helpful to plot the test means for the Types of instruction. These are shown in Figure 1. Here the largest difference appears to be between the didactic approach on the one hand and the empirical and programmed approaches on the other hand. Students in the didactic approach have the lowest scores on the first test. However, they catch up by the third testing and surpass the other two approaches on the fourth testing. Over the five testings, however, these differences cancel, so that there is no overall difference on the average of the five tests between the three Types.

Greater insight into this difference can be seen by examining the ANOVAs for a polynomial contrast transformation. These are shown below

F statistic for Polynomial Contrasts:

<u>Effect</u>	<u>Linear</u>	<u>Quadratic</u>	<u>Cubic</u>	<u>Quartic</u>
Mean	10.16*	0.24	0.01	0.20
Mode	0.07	0.12	1.29	0.97
Type	20.75*	5.15*	1.74	0.75
Mode*Type	0.87	0.08	0.13	1.99
Age	0.04	0.00	0.00	0.17

* $p < .01$

Because the test for sphericity was not significant, all four of these tests are independent. The row for Mean tests whether the means on the five tests increase linearly, quadratically, etc. over time. The mean here is the mean for the entire sample. There is only a linear trend over time for the means. There is no significant effects for Mode, Mode*Type, or Age. The significant linear and quadratic effect for Type suggests that the difference between the didactic and the other approaches is real (see Figure 1). Furthermore, the didactic approach asymptotes at test 4.

Therefore, the deflection in the straight line from Time 4 to Time 5 is real compared to the lack of deflection in either the empirical or programmed approaches.

Thus, at the end of a semester, students appear to achieve the same levels of foreign language mastery regardless of the Mode or Type of teaching. However, the *approach* to this mastery depends upon the Type of instruction. Empirical and programmed learning strategies tend to linearly increase mastery over time. A didactic approach, on the other hand, achieves mastery sooner than the other two, but is more difficult at the early stages of learning.

Repeated Measures / Within Subjects Designs: A quick & dirty approach

Background: There are probably as many different ways to perform repeated measures analysis as there are roads that lead to Rome. Furthermore, there are just as many differences in terminology. Here the term "repeated measures" is used synonymously with "within subjects." Thus, within subjects factors are the same as repeated measures factors. Also note that the SAS use of a "contrast" transformation for repeated measures is not the same as contrast coding as taught by Chick and Gary. Here, the term "transformation" is used to refer to the creation of new dependent variables from the old dependent variables. [Sorry about all this but I did not make up the rules.] The following is one quick and dirty way to perform a repeated measures ANOVA (or regression). There are several other ways to accomplish the same task, so there is no "right" or "wrong" way as long as the correct model is entered and the correct statistics interpreted.

Setting up the data and the SAS commands

1. Make certain the data are entered so that each row of the data matrix is an independent observation. That is, if Abernathy is the first person, belongs to group 1, and has three scores over time (11, 12, and 13). Then enter

Abernathy	1	11	12	13
-----------	---	----	----	----

and not

Abernathy	1	1	11
Abernathy	1	2	12
Abernathy	1	3	13

It is possible to do a repeated measures analysis with the same person entered as many times as there are repeats of the measures, but that type of analysis will not be explicated here.

2. Use GLM and use the model statement as if you were doing a MANOVA. All repeated measures variables are the dependent variables. Suppose the three scores are called SCORE1, SCORE2, and SCORE3 in the SAS data set and GROUP is the independent variable. Then use

```
PROC GLM; CLASSES GROUP;
MODEL SCORE1 SCORE2 SCORE3 = GROUP;
```

3. Use the REPEATED statement to indicate that the dependent variables are repeated measures of the same construct or, if you prefer the other terminology, within subjects factors. A recommended statement is

```
REPEATED <name> <number of levels> <transformation> / PRINTM
PRINTE SUMMARY;
```

where <name> is a name for the measures (or within subject factors), <number of levels> gives the number of levels for the factor, and <transformation> is the type of multivariate transformation. For our example,

REPEATED TIME 3 POLYNOMIAL / PRINTM PRINTE SUMMARY;

will work just fine.

When there is more than a single repeated measures factor, then you must specify them in the correct order. For example, suppose the design called for a comparison of recall versus recognition memory for phrases that are syntactically easy, moderate, and hard to remember. Each subject has $2 \times 3 = 6$ scores. Suppose Abernathy's scores are arranged in the following way:

		Recall			Recognition		
		Easy	Mod	Hard	Easy	Mod	Hard
Group		Y1	Y2	Y3	Y4	Y5	Y6
Abernathy	1	12	8	3	21	16	14

The SAS statements should be:

```
PROC GLM; CLASSES GROUP;
  MODEL Y1-Y6 = GROUP;
  REPEATED MEMTYPE 2, DIFFCLTY 3 POLYNOMIAL / PRINTM PRINTE
SUMMARY;
```

There we specify two repeated measures factors (or within subjects factors). The first is MEMTYPE for recall versus recognition memory, and the second is DIFFCLTY and to denote the difficulty level of the phrases. . **Note that the factor that changes least rapidly always comes first.** Had we specified DIFFCLTY 3, MEMTYPE 2, then SAS would have interpreted Y1 as Recall-Easy, Y2 as Recognition-Easy, Y3 as Recall-Moderate, etc.

4. Remember that using a REPEATED statement will always generate a transformation of the variables. Always choose the type of transformation that will reveal the most meaningful information about your data.

5. That is all there is to doing a repeated measures ANOVA or Regression. You can use the CONTRAST statement if you wish to contrast code categorical independent variables. Just make certain that you place the CONTRAST statement before the REPEATED statement.

Interpreting the Output

This is a synopsis of the handout on *Repeated Measures*. You should follow these steps to interpret the output.

1. The first thing SAS writes in the output is the design for the repeated measures. Always check this to make certain that correctly specified the levels of the repeated measures. This is particularly important when there is more than a single within subjects factor.
2. The second thing to check is whether error covariance matrix can be orthogonally transformed. The tests of sphericity will tell you that. Some transformations in SAS are deliberately set up to be orthogonal (e.g., POLYNOMIAL with no further qualifiers); other transformations are not orthogonal (e.g., CONTRAST). If a transformation is orthogonal, then SAS will print out one test of sphericity. If a transformation is not orthogonal, then SAS spits out two tests of sphericity. The first test is for the straight transformation. The second test is for the orthogonal components of the transformation. In this case, it is the second test--the one for the orthogonal components--that you want to interpret.
3. If the c^2 test for sphericity is not significant, then ignore all the MANOVA output and interpret the RM ANOVA results for the within subjects effects. These are labelled in the SAS output as "Univariate Tests of Hypotheses for Within Subjects Effects."
4. If the c^2 test for sphericity was significant, then you can interpret the MANOVA results or the adjusted probability levels from Greenhouse-Geisser and the Huynh-Feldt corrections for the within subjects effects. It is often a good idea to compare the MANOVA significance with the Greenhouse-Geisser and the Huynh-Feldt adjusted significance levels to make certain there is agreement between them.
5. The between subjects effects are not affected by the results of the sphericity test. Hence, SAS output with the heading "Tests of Hypotheses for Between Subjects Effects" will always be correct.

6. Always interpret the output for the transformed variables. It can often tell you something important about the data. Exactly what it tells you will depend upon the type of transformation you used in the REPEATED statement.
7. Always make certain that the raw means and standard deviations are printed. If you have not gotten them in the GLM procedure with the MEANS statement, then get them by using PROC MEANS, PROC UNIVARIATE, or PROC SUMMARY. Repeated measures or within subjects designs are useless when the results are not interpreted with respect to the raw data.